

MATH 6642 – Longitudinal Data Analysis

Sample Exam Questions

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Useful formulas:

- $(A + UDV)^{-1} = A^{-1} - A^{-1}U(D^{-1} + VA^{-1}U)^{-1}VA^{-1}$
- 1. Write a brief discussion with illustrations (using ellipses in beta space if you wish) explaining the following issues concerning the interpretation of p-values in the output produced by the ‘summary’ method applied to a mixed model for longitudinal data analysis.
 - a) [5] Why should p-values for main effects that are marginal to an interaction be interpreted with caution. Are there any situation in which it is legitimate to use them?
 - b) [5] If two main effects that are not marginal to any interaction have non-significant p-values, is it appropriate to drop both terms?
 - c) [5] If a time-varying variable is included as a raw variable and as a contextual variable, what is the possible consequence of dropping the contextual variable? When would dropping it result in no change in the estimated coefficient of the raw variable?
- 2. [10] Consider a random coefficients model for a model with time as a single predictor with a random intercept and a random slope. Let the G matrix have the decomposition $G = AA'$ where

$$A = \begin{bmatrix} a_{00} & 0 \\ a_{10} & a_{11} \end{bmatrix}$$

is lower triangular. Explain and justify the interpretation of the elements of A in terms of the properties of the random true regression lines of Y on T as they vary from subject to subject. Draw a sketch to illustrate your answer.

- 3. Consider a longitudinal random slopes model where the model for i th subject is given by

$$Y_i = X_i\gamma + X_iu_i + \epsilon_i$$

with X_i a $T_i \times 2$ full column rank matrix with a first column of 1's and a second column containing the times at which the vector of responses Y_i was observed, $u_i \sim N_2(0, G)$, with G positive definite, independent of $\epsilon_i \sim N(0, \sigma^2 I)$.

Let $\beta_i = \gamma + u_i$. In the following questions, treat the macro level parameters γ, G, σ^2 as if they were as they were fixed and known.

- a) [5] Give an expression for the BLUE, $\hat{\beta}_i$, of β_i .
- b) [5] Give an expression for the BLUP, $\tilde{\beta}_i$, of β_i as a function of the BLUE.
- c) [5] Describe how the BLUP is a shrinkage estimator based on the BLUE.
- d) [5] For large T_i and relatively fixed variance for the values of time, how will the BLUP behave relative to the BLUE?
- e) [5] What are the implications for the BLUP if G is highly concentrated?
- f) [10] Show that $Var(\tilde{\beta}_i) < G < Var(\hat{\beta}_i)$. Note: for two symmetric matrices, A and B , $A < B$ if and only iff $B - A$ is positive definite.
- 4. The following questions refer to the output on the last page for a mixed model fitted on the full high school math achievement data set. The model uses SES, SES.School which is the mean SES in the sample in each school, ‘female’ which is an individual level indicator variable and ‘Type’ which is three-level factor with levels “Coed”, “Girl” and “Boy” with the obvious definition and Sector which is a 2-level factor with levels “Catholic” and “Public”.
 - a. [5] Consider two Public coed schools, one with mean SES = 0 and the other with mean SES = 1. Suppose the values of SES in the former school range from -1 to 1 and in the latter school from 0 to 2. Draw a graph showing the predicted MathAch in these two schools over the range of values

- of SES in each school. On your graph identify the value and location of the contextual effect of SES, the within school effect of SES and the compositional effect of SES.
- b. [5] Suppose you were to refit the model without SES.School. What would you expect would happen to the coefficient for SES? Would it stay roughly the same, get bigger or get smaller, or is the change unpredictable? Explain.
 - c. [5] Suppose you want to perform an overall test of the importance of gender in the model, either within schools or between schools. Specify a hypothesis matrix that would perform this test.
 - d. [5] Estimate the difference between the predicted math achievement of a boy in a boys' school versus a girl in a girls'school. If you suspected that this is affected by the SES of the school, how would you modify the model to test this hypothesis?
5. [10] Consider a mixed model for data with a response variable Y and regression on an intercept and a single predictor X . Suppose that each cluster has size 3 and X is observed twice at the value 0 and once at the value 1; i.e. the observed values of X are 0, 0, 1 within each cluster. Show whether it is possible to identify the variance parameters of a random slope model?
 6. Data is obtained on the effects of two treatments (A and B). Data was recorded on symptoms weekly throughout 10 weeks of active treatment and then for 8 more weeks following termination of the treatment. A linear spline model is fit to the data and the following results were obtained:

```

sp <- function( x ) {
  gsp( x,
        knots = c( 10 ),      # 1 knots => 2 intervals
        degree = c( 1 , 1 ) , # linear in each interval
        smooth = c( 0 )       # continuous at the knot
      )
}

> fit <- lme( Symptom ~ sp(Weeks)*tx, random=~1+Weeks|id, data = d)
> summary(fit)
Linear mixed-effects model fit by REML
Data: d
      AIC      BIC    logLik
15099.44 15153.18 -7539.719

Random effects:
Formula: ~1 + Weeks | id
Structure: General positive-definite, Log-Cholesky parametrization
          StdDev   Corr
(Intercept) 37.921194 (Intr)
Weeks        2.395329 -0.218
Residual    23.285420

Fixed effects: Symptom ~ sp(Weeks) * tx
              Value Std.Error   DF   t-value p-value
(Intercept)   131.84484   6.450470 1516   20.43957  0.0000
sp(Weeks)D1(0) -11.21889   0.508022 1516  -22.08346  0.0000
sp(Weeks)C(10).1  18.24584   0.571104 1516   31.94835  0.0000
txB            -2.61621   9.122343   78   -0.28679  0.7750
sp(Weeks)D1(0):txB  1.55017   0.718452 1516    2.15765  0.0311
sp(Weeks)C(10).1:txB -4.83791   0.807664 1516   -5.99001  0.0000

. . .

Standardized Within-Group Residuals:
          Min           Q1           Med           Q3           Max
-4.109811205 -0.593593851 -0.002546668  0.654503092  3.122385555

Number of Observations: 1600
Number of Groups: 80

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- a) [5] Interpret each fixed effects coefficients in the model drawing an appropriate graph.
- b) [5] In a separate graph, sketch the predicted trajectory for $\mathbf{tx}$ equal to A.
7. [5] What is Simpson's Paradox and how is it related to the inclusion of contextual variables in hierarchical models?
8. [5] In the model used in question 5, why would one want to add the average SES of each school, 'SES.School', to the model?
9. [5] Let Σ be symmetric. Show that Σ is positive-definite if and only there exists a non-singular matrix A such that $\Sigma = AA'$.

Output for question 4

Linear mixed-effects model fit by REML

Data: hs

	AIC	BIC	logLik
	46477.96	46581.14	-23223.98

Random effects:

Formula: ~1 + SES + female | School

Structure: General positive-definite, Log-Cholesky parametrization

	StdDev	Corr
(Intercept)	1.6500478	(Intr) SES
SES	0.2661823	0.755
female	0.9507088	-0.606 -0.979
Residual	6.0343349	

Fixed effects: MathAch ~ SES * Sector + SES.School + female + Type

	Value	Std.Error	DF	t-value	p-value
(Intercept)	13.633578	0.3315175	7022	41.12476	0.0000
SES	1.436456	0.1620660	7022	8.86340	0.0000
SectorPublic	-0.888000	0.3711059	155	-2.39285	0.0179
SES.School	3.082877	0.3662818	155	8.41668	0.0000
female	-1.078197	0.1902275	7022	-5.66793	0.0000
TypeGirl	0.258757	0.4770260	155	0.54244	0.5883
TypeBoy	1.358493	0.5415658	155	2.50845	0.0132
SES:SectorPublic	1.283763	0.2121825	7022	6.05028	0.0000

Correlation:

	(Intr) SES	SctrPb	SES.Sc	female	TypGrl	TypeBy
SES	-0.010					
SectorPublic	-0.807	-0.010				
SES.School	-0.273	-0.199	0.374			
female	-0.381	-0.023	0.020	0.006		
TypeGirl	-0.550	0.037	0.561	0.206	-0.134	
TypeBoy	-0.591	0.036	0.465	0.088	0.233	0.321
SES:SectorPublic	0.056	-0.733	0.012	-0.002	-0.001	-0.058 -0.045

Standardized Within-Group Residuals:

	Min	Q1	Med	Q3	Max
	-3.10239407	-0.73144171	0.02590791	0.75673178	2.99519661

Number of Observations: 7185

Number of Groups: 160