

Monte-Carlo Markov Chain

When i.i.d Monte-Carlo is too hard.

Can't sample from f directly on sample space S

But there is a conditional distribution

$$q_f(x_{t+1} | x_t)$$

So starting at some x_0

and adding $x_1 \sim q_f(\cdot | x_0)$

⋮

$$x_{t+1} \sim q(\cdot | x_t)$$

⋮

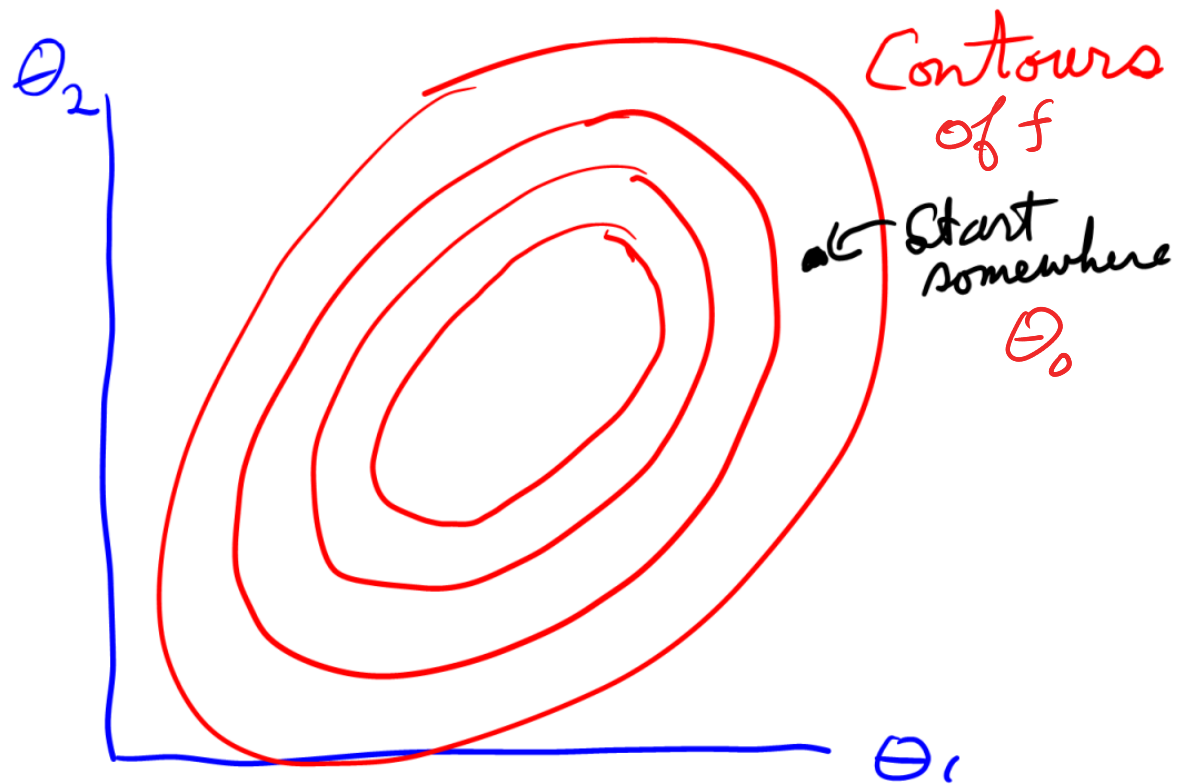
creates a sequence $x_t, \dots, x_t, x_{t+1}, \dots, x_{t+1}$
that is $\sim f$ but not i.i.d.

Questions:

- How to make this work?
- When will it work?

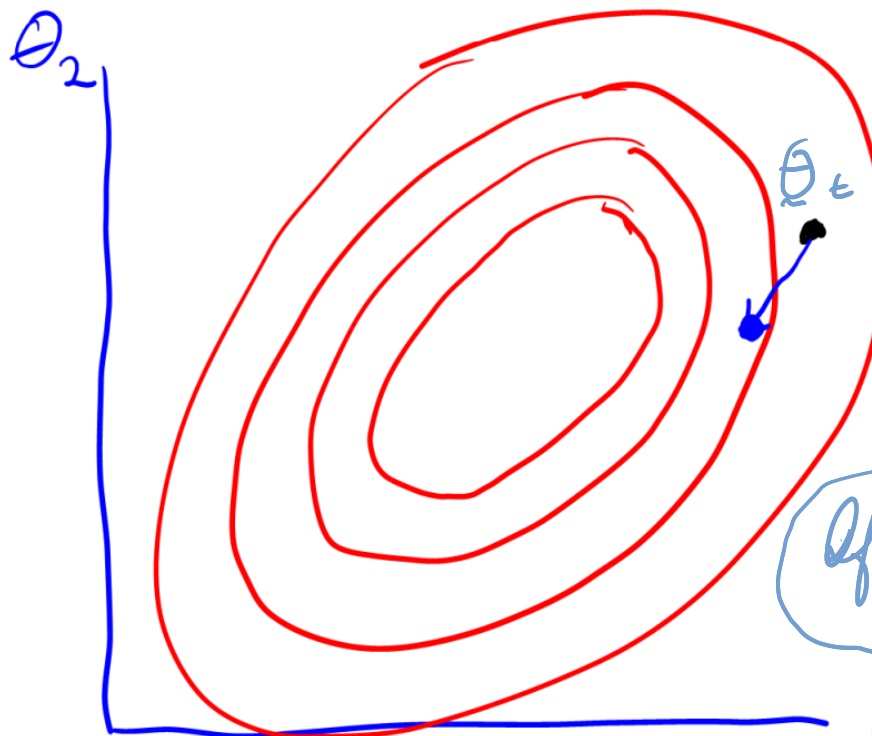
One way to make it work:

Metropolis-Hastings algorithm.



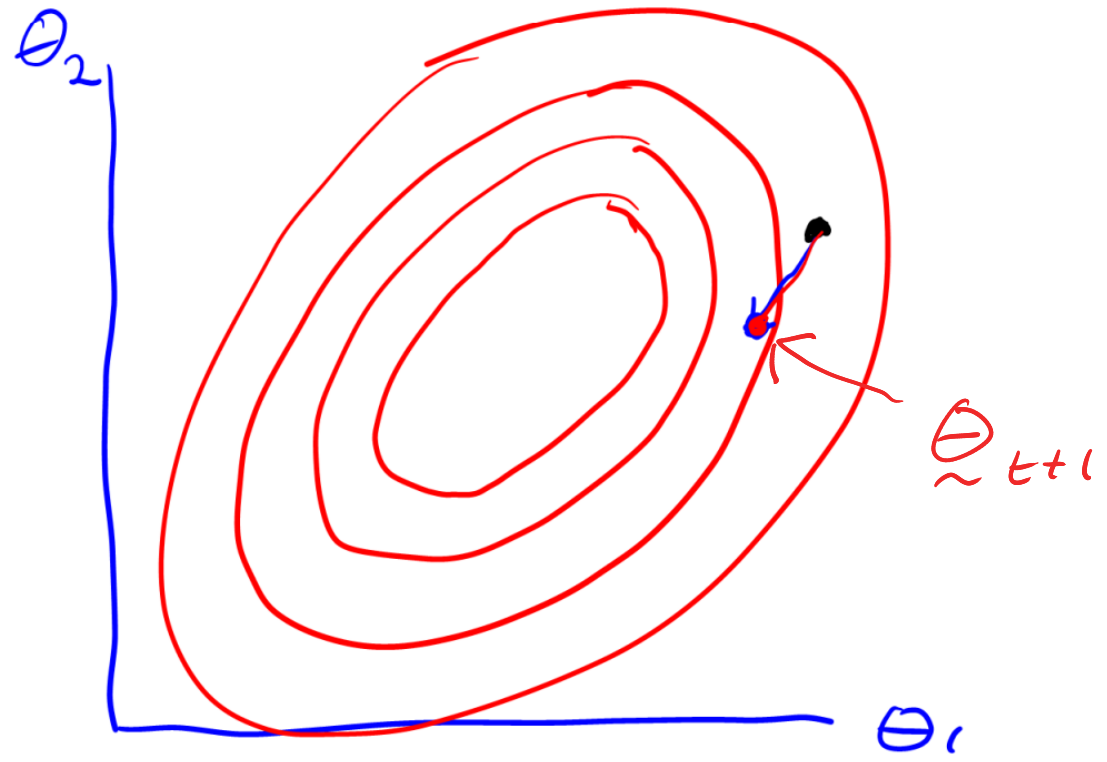
Using θ
instead
of μ
but same
idea.

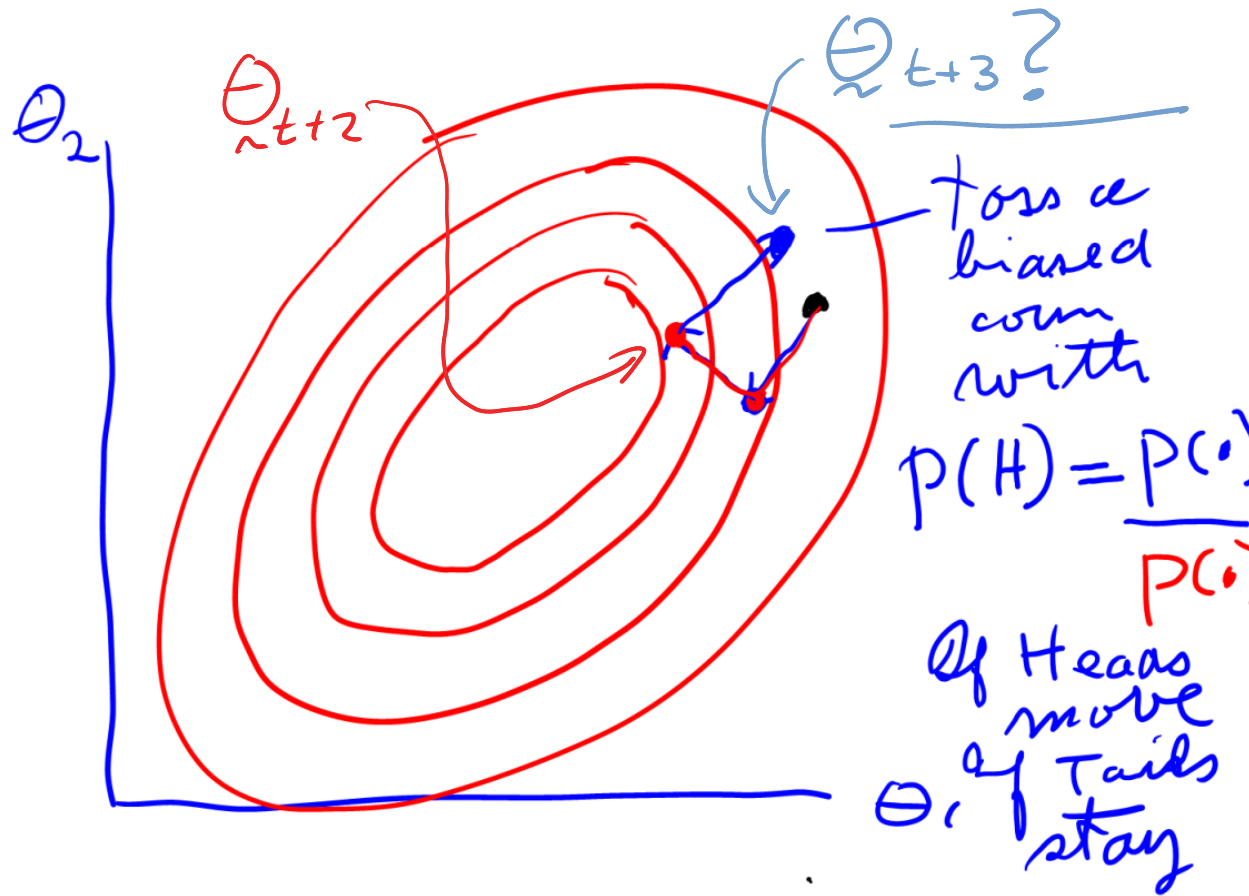
Metropolis-Hastings algorithm
actually this is the special case.



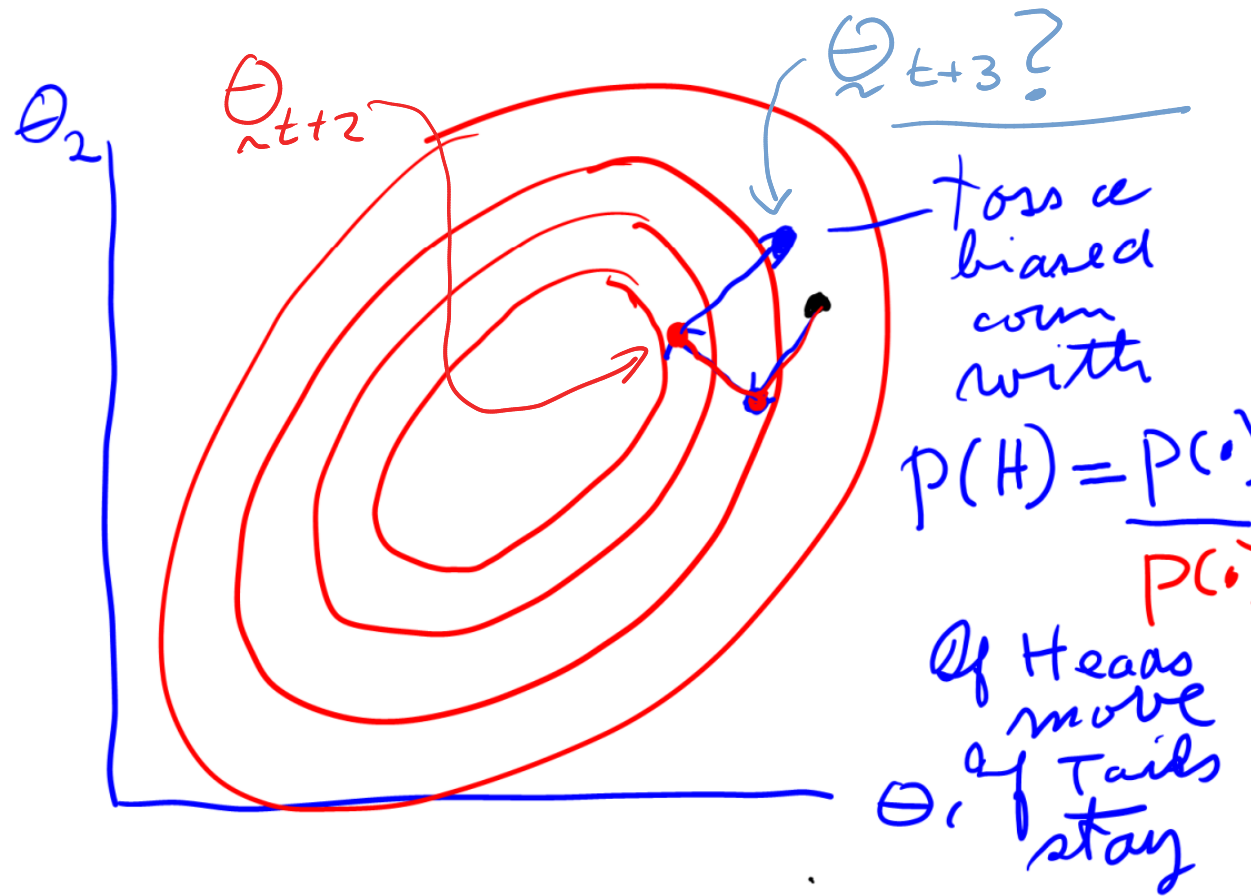
Choose a
random
step.

If $\frac{P(\bullet)}{P(\circ)} > 1$
 θ_t then
 $\theta_{t+1} = \bullet$



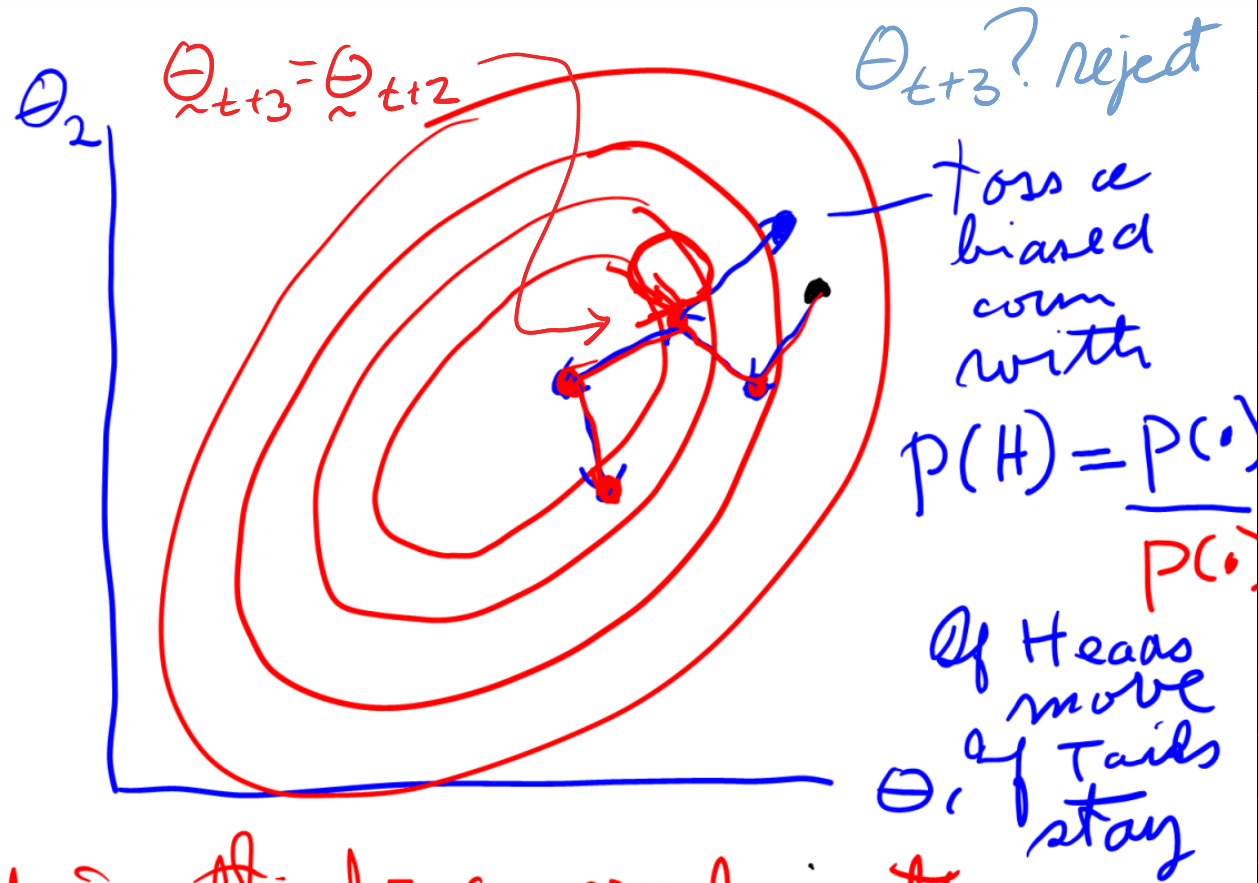


This is a very big deal! Why?



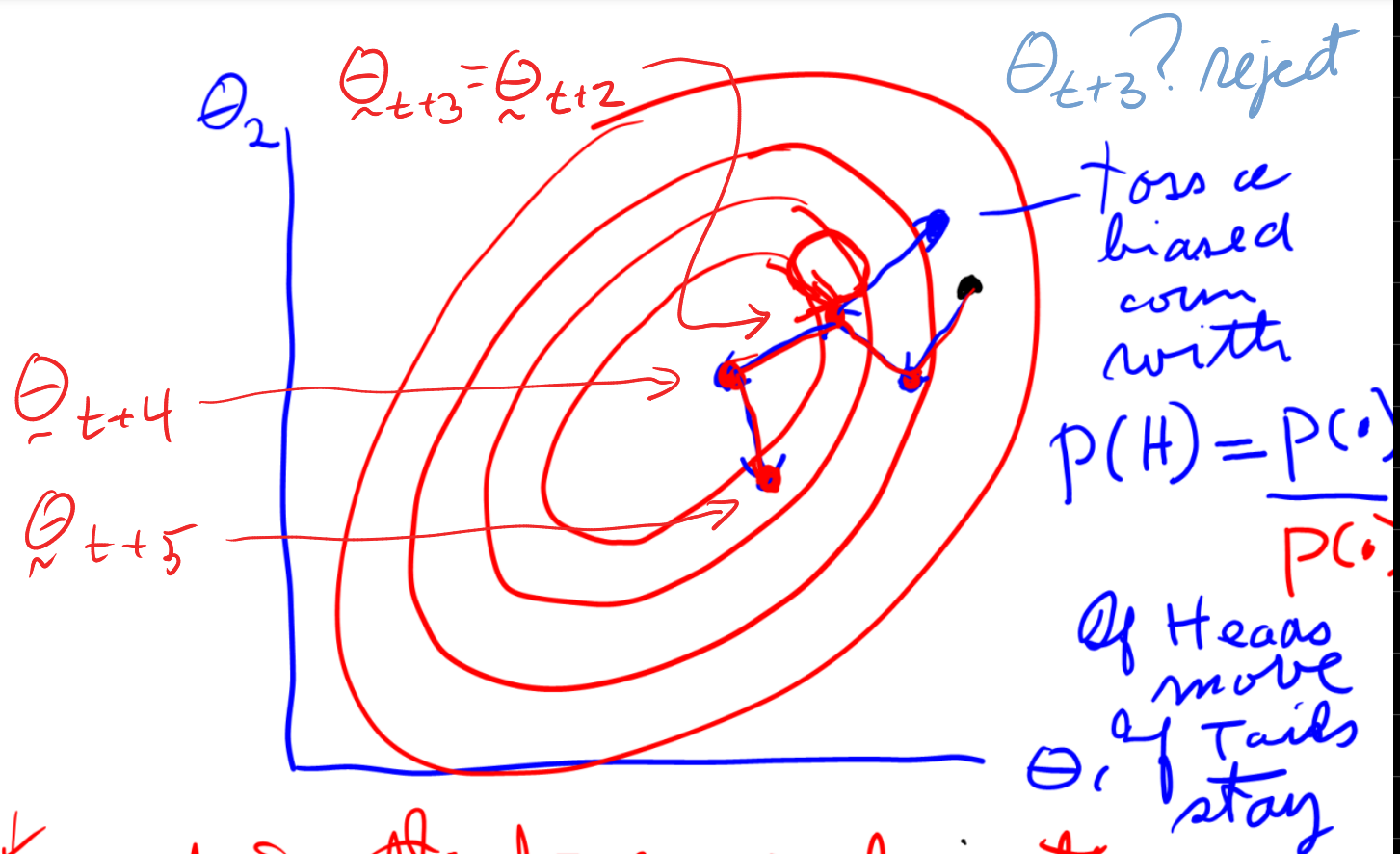
This is a very big deal!

We don't need p itself!! Any $f \propto p$ will do.



Keep doing this for a very long time

Being able to use $f \propto p$ was revolutionary for Bayesian stats.



Keep doing this for a very long time

We get a random chain.

$x_0, x_1, x_2, \dots, x_T, \dots, x_t, x_{t+1}, \dots, x_N$

Problems: (1) early part influenced by choice of x_0
(2) not i.i.d.

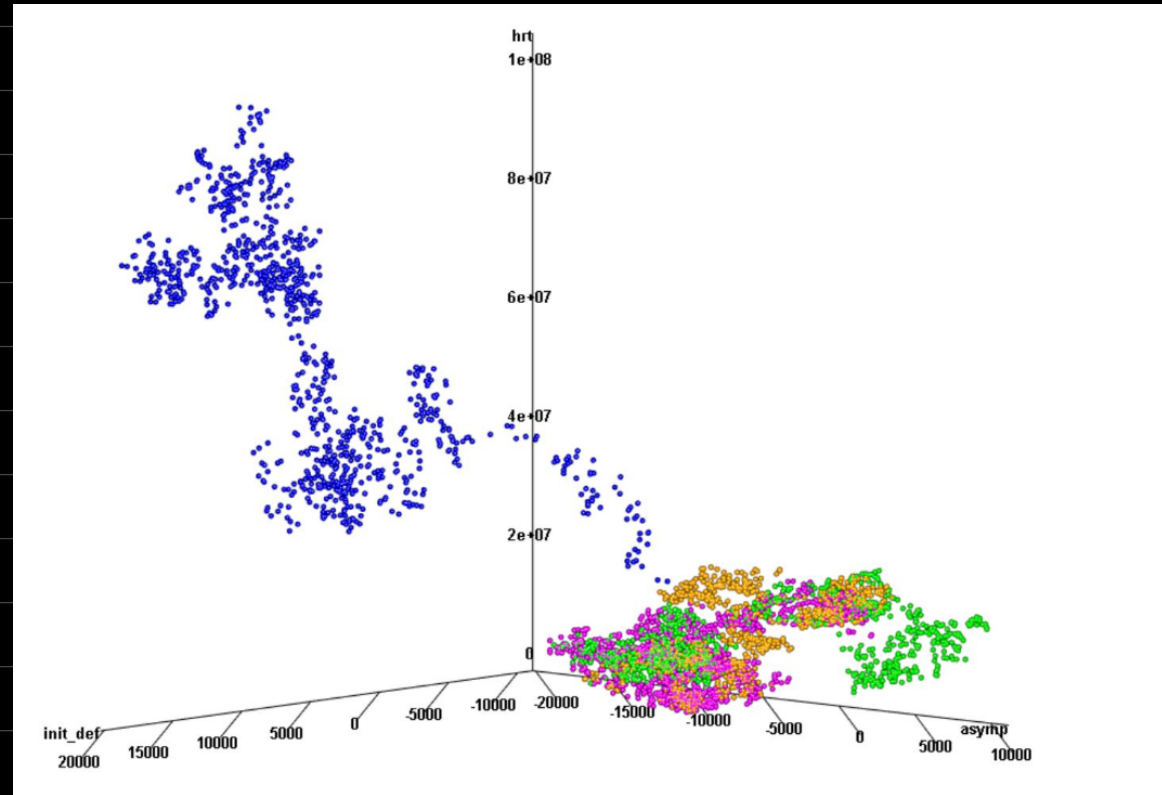
Some parts of f might have been ignored.

Ideas? What could we do?

- (1) Throw out early part? "Burn in"
- (2) Thin the chain: Run longer & keep every 10^{th} $\{x_t\}$

(3) Create many chains with different starting values and see if they converge.

Diverging chains:



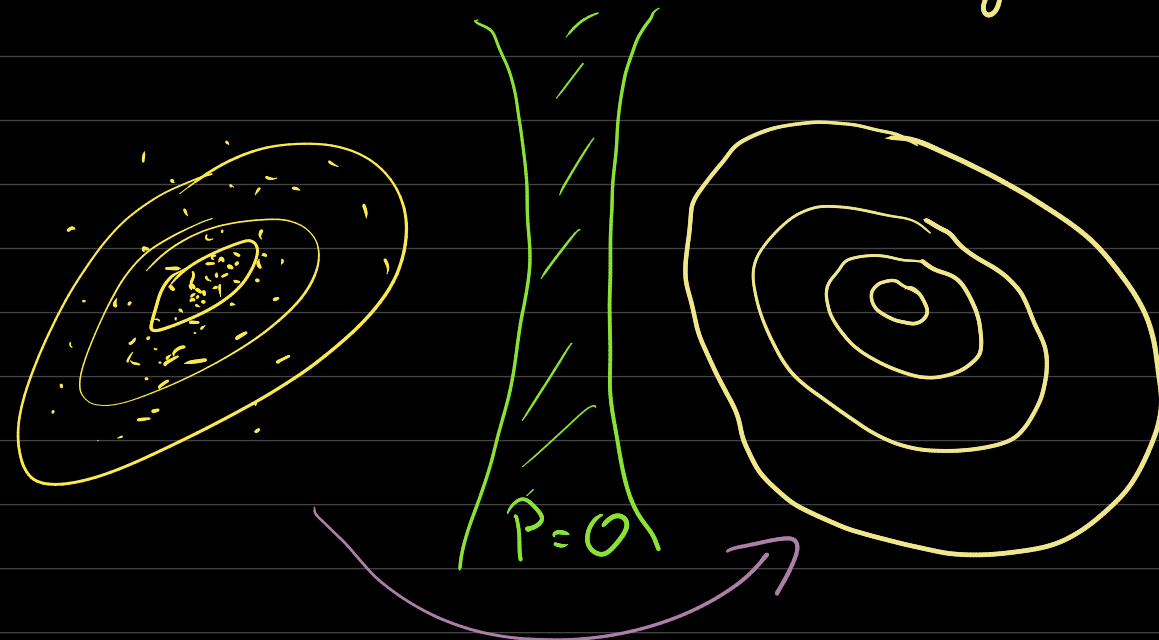
But first:

- When will MCMC utterly fail?
- Travelling around a mountain range
- you want to step on every part in a way that is proportional to altitude (f)



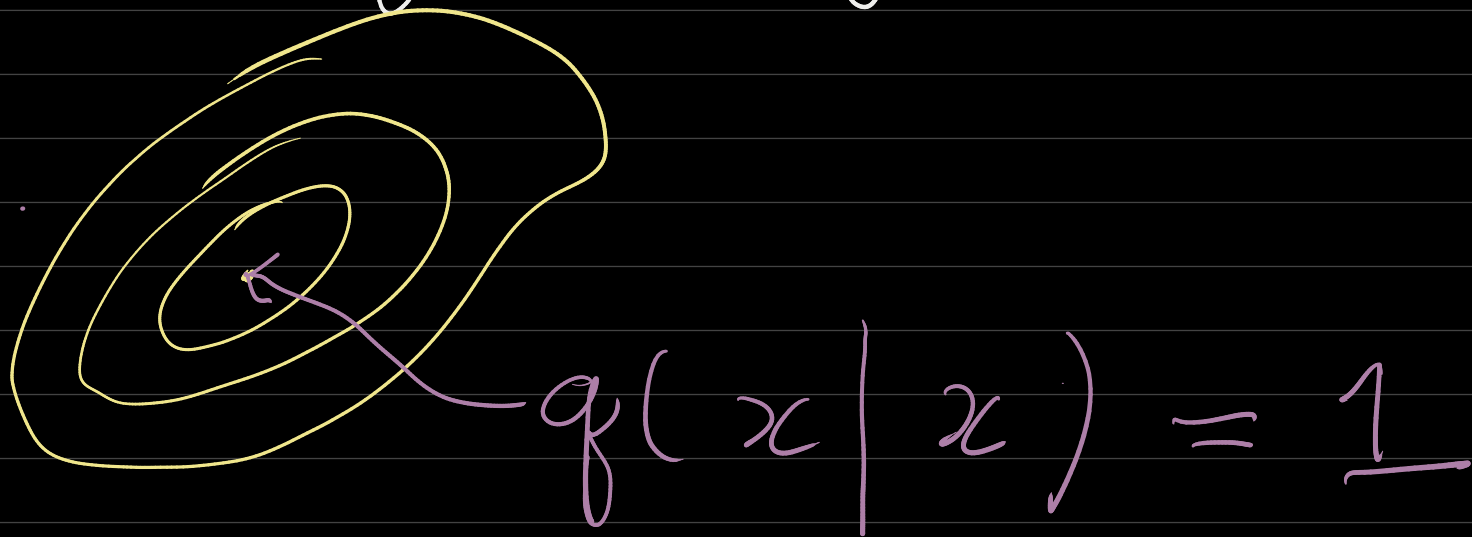
What can go wrong?

- ① Parts of range separated by uncrossable chasm (region of 0 cond'l prob.)



MCMC will never reach peaks where it did not start.

② Peaks you can't get off of.



Theory of MCMC addresses this:

Discrete finite sample space case

$$X_0, X_1, \dots, X_t, X_{t+1}, \dots$$

with $X_t \in \mathcal{S}$, state space $\mathcal{S} = \{1, \dots, n\}$

Possible states
↓

So $q(x_{t+1} | x_t)$ can be represented
in a transition probability matrix

Markov
because ...?

$$P = [P_{ij}], \quad P_{ij} = P_n(X_{t+1} = j | X_t = i)$$

E.g. $\mathcal{S} = \{1, 2, 3\}$

$$P = \begin{bmatrix} 0.1 & 0.5 & 0.4 \\ 0.2 & 0.0 & 0.8 \\ 0.3 & 0.0 & 0.7 \end{bmatrix}$$

rows sum to 1

i th row is
conditional probs
for $X_{t+1} | X_t = j$

Suppose $x_0 = 2$

If X_t has pmf $\underline{\pi}_t$ (vector of probabilities)
then X_{t+1} has pmf $\underline{\pi}_t P$ (work it out)
 X_{t+2} has pmf $\underline{\pi}_t P^2$
 \vdots
 X_{t+n} has pmf $\underline{\pi}_t P^n$

Properties:

$$P \underline{1} = \underline{1}$$

Row sums = 1

$\underline{1}$ is a right-eigenvector
belonging to eigenvalue 1

Definition: P is irreducible

- iff any state can eventually be reached from any other state
- iff $\exists m \in \mathbb{Z}^+$ all elements of P^m are positive.

Example: $\begin{bmatrix} 1 & 0 & 0 \\ 0 & 0.5 & 0.5 \\ 0 & 0.5 & 0.5 \end{bmatrix}$ not irreducible
"is reducible"

Less obvious: $P = \begin{bmatrix} 1 & 0 & 0 \\ 0.5 & 0 & 0.5 \\ 0 & 0.5 & 0.5 \end{bmatrix}$

is not irreducible. $\{1\}$ is an absorbing state.